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ANDREWS' PLOTS AND THEIR APPLICATIONS

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ABSTRACT

A brief survey of several graphical multivariate techniques are given.

Andrews' method is exploited as a graphical tool for the examination of changes over time in the parameters of a time series model. An example is given to illustrate the method

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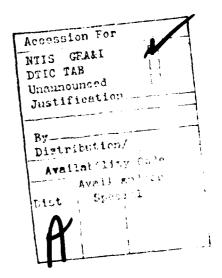
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SIGNIFICANCE AND EXPLANATION

A brief survey of several graphical multivariate techniques are given.

One of these due to Andrews is given in more detail. In his method, Andrews represents each multidimensional point by a Fourier function. The clustering of plots of these functions is equivalent to the clustering of the multidimensional points. Andrews' method is exploited as a graphical tool for the examination of changes over time in the parameters of a time series model. An example consisting of temperature data is given to illustrate the method.



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ANDREWS' PLOTS AND THEIR APPLICATIONS

Agnes M. Herzberg

1. Introduction.

If multivariate data are m-dimensional, then each set of m measurements can be represented as an m-dimensional point. For m = 1,2, these points may be plotted and clusters easily determined by inspection. For m > 2, this is more difficult. Several authors have developed graphical techniques to plot high-dimensional data in two dimensions in order to be able to visually cluster the data; see, for example, Andrews (1972), Chernoff (1973), Kleiner and Hartigan (1981) and Anderson (1928, 1936). More mathematical techniques have been given by Beale (1969) and Banfield and Bassil (1977).

2. Several graphical methods.

Kleiner and Hartigan (1981) introduced what they termed trees and castles. First, a hierarchical clustering algorithm is applied to the m variables over all the points; see for example Gnanadesikan (1977). From this the structure of the tree or castle will be determined. All points will be represented by a similar structure, i.e. the thickness, position and angle of the branches in the case of trees will be the same, but the length of the branches will be determined by the sizes of the respective variables for the individual points. Similar trees and castles determined by visual inspection are clustered.

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Chernoff (1973) represents each variable as a feature of a face, for example length of mouth, shape of mouth, size of eyes, etc. The resulting clustering of this representation is very subjective because different people focus on different features of faces.

Anderson (1928, 1936) was very concerned with the sepal length and width and petal length and width of irises. He developed pictorial methods which he called ideographs for representing and comparing these four-dimensional data. An ideograph looks like an upside-down U with some width. In the case of the iris measurements, the inside and outside height and width measurements of the ideograph are proportional to the sepal length and width and the petal length and width, respectively. Similar ideographs can easily be clustered by visual inspection.

There are many other graphical representations for multivariate data, but these will not be discussed.

3. Andrews' plots.

Andrews (1972) proposed the following simple and useful method of plotting high-dimensional data in two dimensions. If the data are m-dimensional, each point $\mathbf{x}' = (\mathbf{x}_1, \dots, \mathbf{x}_m)$, where $\mathbf{x}_i (i = 1, \dots, m)$ are the measured variables, is represented by the function

$$f_x(t) = x_1 2^{-1/2} + x_2 \sin t + x_3 \cos t + x_4 \sin 2t + x_5 \cos 2t + \cdots$$
 (1)

plotted over the range $-\pi < t < \pi$. The functions given by (1) have several properties. If $x_i = (x_{1i}, \cdots, x_{mi})$ (i = 1, ..., n) are n points in m-dimensional space, then

$$f_{\underline{x}}(t) = \frac{1}{n} \sum_{i=1}^{n} f_{\underline{x}_{i}}(t)$$

$$\|f_{\underline{x}_{i}}(t) - f_{\underline{x}_{j}}(t)\|_{L_{2}} = \int_{-\pi}^{\pi} \{f_{\underline{x}_{i}}(t) - f_{\underline{x}_{j}}(t)\}^{2} dt$$

$$= \pi \|\underline{x}_{i} - \underline{x}_{j}\|^{2}$$

$$= \pi \sum_{k=1}^{m} (x_{ki} - x_{kj})^{2}.$$

Thus Andrews' plots will preserve means, distances and variances and will also give one-dimensional projections. When (1) is plotted for each data point \times , the clustering of the points may be seen by a banding together of the plots of the functions. Since the functions preserve the distance property, plots of the functions that are close together imply that the corresponding data points are close together.

4. Variation of model parameters.

Herzberg and Hickie (1981) considered the following. Let the regression model be written in the form

$$Y_{\sim j} = X \beta_j + U (j = 1, \cdots, T-n+1), \qquad (2)$$

where T is the total number of observations n is the number of observations in each subgroup of observations used for estimating the unknown parameters, $Y_j = (y_{1j}, \cdots, y_{nj})^*$ is an $n \times 1$ vector, y_{kj} being the k^{th} observation in the j^{th} subgroup $(k = 1, \cdots, n)$, X is the $n \times m$ matrix of the regressors, β_j is the $m \times 1$ vector of unknown parameters and Y_j is the $n \times 1$ vector of error terms. All the elements of the Y_j 's are assumed to be independent and normally distributed with mean 0 and variance y_j . It is assumed that the T observations are taken sequentially over time and it is desired to examine the variations in the y_j over time.

Let $\hat{\beta}_j = (\hat{\beta}_{1j}, \cdots, \hat{\beta}_{mj})^*$ be the m × 1 vector of least squares estimates of the elements of the vector $\hat{\beta}_j$ obtained from the jth set of n observations (n < T), i.e. $\hat{\beta}_1$ is estimated from the first n observations, $\hat{\beta}_2$ is estimated from the second observation to the (n+1)st observations, etc. From each $\hat{\beta}_j$ plot the function $\hat{\beta}_j$ (t), defined in (1), over the range $-\pi < t < \pi$. The plots of these functions will show the change over time in the vector of coefficients $\hat{\beta}_j$.

Herzberg and Hickie (1981) consider two sets of data using polynomial and Fourier series models in (2). One of the sets of data has a cyclic effect, the other having cyclic effect plus trend. For both sets of data it was known that the period was 12 months. It could also be seen that every 12th plot was similar.

For one of their examples, namely the monthly mean daily air temperatures (°C) at sea level for England and Wales from January 1970 to December 1977 as published by the Central Statistical Office Monthly Digest of Statistics (HMSO), Herzberg and Hickie (1981) fitted the cubic polynomial model,

$$E(y_{j+i-1}) = \beta_{1j} + \beta_{2j}i + \beta_{3j}i^2 + \beta_{4j}i^3,$$
 (3)

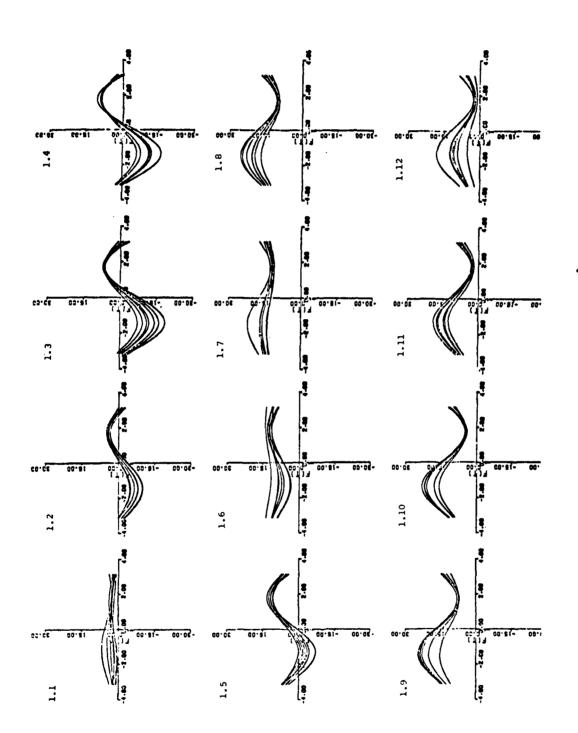
by least squares. Here y_{j+i-1} is the observed temperature in mouth j+i-1 For each j in (3) fixed, $i=1,\cdots,12$, $\hat{\beta}_j=(\hat{\beta}_{1j},\hat{\beta}_{2j},\hat{\beta}_{3j},\hat{\beta}_{4j})$ the least squares estimate of β_j was obtained $(j=1,\cdots,85)$. Figure 1 shows the resulting Andrews' plots. The plots in Figure 1.k are those obtained from $\hat{\beta}_j$ $(j=k,\ k+12,\ k+24,\ k+48,\ k+60,\ k+72,\ k+84;\ j<85)$. It can be seen that the plots in each of the Figures 1.k are similar.

In situations where the period is unknown, Andrews' plots may be plotted for several values of n in order to determine similarities and, therefore, the length of the period.

Because of their mathematical and resulting statistical properties,

Andrews' plots can be used as a tool for finding outliers in a time series.

Work is at present being done on this and on using Andrews' plots as a
sequential graphical method for discriminating among models.



January 1970 to December 1977; Figure 1.k (k = 1, ..., 12) consists of Andrews' plots at sea level for England and Wales, for j = k, k + 12, k + 24, k + 36, k + 48, k + 60, k + 72, k + 84 (j < 85). $\hat{\beta}$ obtained from (3) using Figure 1. Andrews' plots $f_{\beta}(t)$ (j = 1, •••, 85), the monthly mean daily air temperature (°C)

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